

Geometry of q -normal distribution
through non-Fisher metric

Masaru Tanaka

Department of Information and Computer Sciences

Saitama University

$$\text{Tsallis entropy} \quad \frac{1}{1-q} \left\{ \int dx p_q(x)^q - 1 \right\}$$

Given two independent systems A and B with Tsallis entropy S_q^A and S_q^B respectively.

Let Tsallis entropy of the compound system $A \cup B$ be $S_q^{A \cup B}$, then

$$S_q^{A \cup B} = S_q^A + S_q^B + (1 - q)S_q^A S_q^B$$



Non-additive entropy

Boltzmann-Shannon entropy ($q = 1$) is the special case, that is additive.

The q-normal distribution is obtained by the following variation problem

Tsallis entropy

The 1st q-moment

$$\begin{aligned} & \frac{1}{1-q} \left\{ \int dx p(x)^q - 1 \right\} \\ & + A \left\{ \int dx p(x) - 1 \right\} + B \left\{ \frac{1}{D} \int dx p(x)^q x - \mu \right\} \\ & + C \left\{ \frac{1}{D} \int dx p(x)^q (x - \mu)^2 - \sigma^2 \right\} \end{aligned}$$

The normalization

The 2nd q-moment

$$D = \int dx p(x)^q$$

The probability distribution has its own entropy.

q-EXPECTATION AND q-NORMAL DISTRIBUTION

$$q\text{-expectation : } E_q[f(x)] = \Omega_q^2 \int dx p(x)^q f(x)$$

➔ 1st q-moment μ , 2nd q-moment σ^2

$$\text{Tsallis entropy : } S_q = \frac{1}{1-q} \left\{ \int dx p(x)^q - 1 \right\}$$

0 < q < 3
q-normal distribution

$$p_q(x) = \frac{1}{Z_q} \left\{ 1 - \frac{1-q}{3-q} \frac{(x-\mu)^2}{\sigma^2} \right\}^{\frac{1}{1-q}} \quad \text{where } Z_q = A_q \sigma \text{ and}$$

$$\left\{ \begin{array}{l} A_q = \frac{\sqrt{\frac{3-q}{q-1}} \pi \Gamma\left(\frac{1}{q-1} - \frac{1}{2}\right)}{\Gamma\left(\frac{1}{q-1}\right)} \quad \text{for } 1 \leq q < 3 \\ \\ A_q = \frac{\sqrt{\frac{3-q}{1-q}} \pi \Gamma\left(\frac{2-q}{1-q}\right)}{\Gamma\left(\frac{2-q}{1-q} + \frac{1}{2}\right)} \quad \text{for } 0 < q < 1 \end{array} \right.$$

$$\text{An escort distribution : } \Omega_q^2 p_q(x)^q = \frac{2}{3-q} \frac{1}{Z_q} \left\{ 1 - \frac{1-q}{3-q} \frac{(x-\mu)^2}{\sigma^2} \right\}^{\frac{q}{1-q}},$$

$$\text{where } \Omega_q^2 = \frac{2}{3-q} Z_q^{q-1} \text{ for } q\text{-normal distribution}$$

- **q-normal distribution**

- for $q = 1$, **the normal distribution**

$$p_1(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\}$$

- for $q = 2$, **Cauchy distribution**

$$p_2(x) = \frac{1}{B\left(\frac{1}{2}, \frac{1}{2}\right)\sigma} \left\{1 + \frac{(x-\mu)^2}{\sigma^2}\right\}^{-1}$$

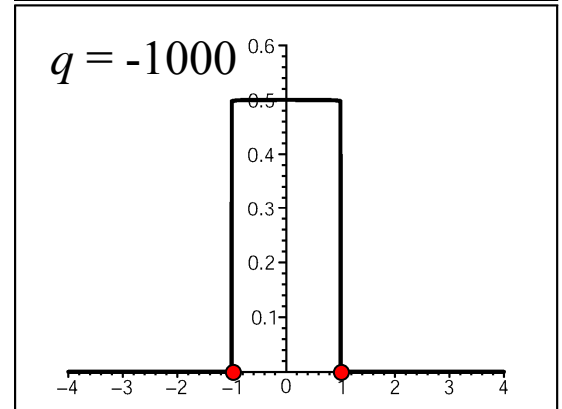
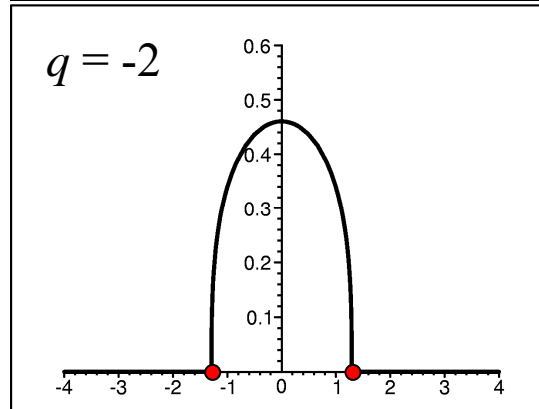
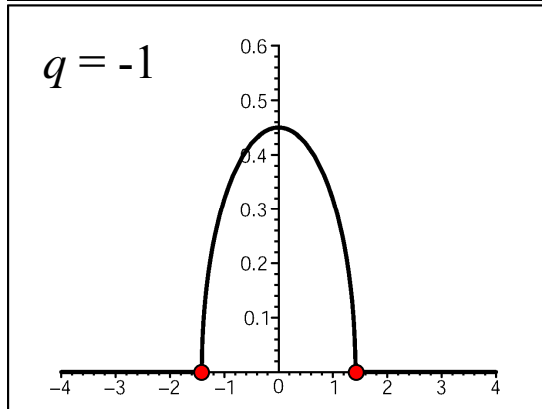
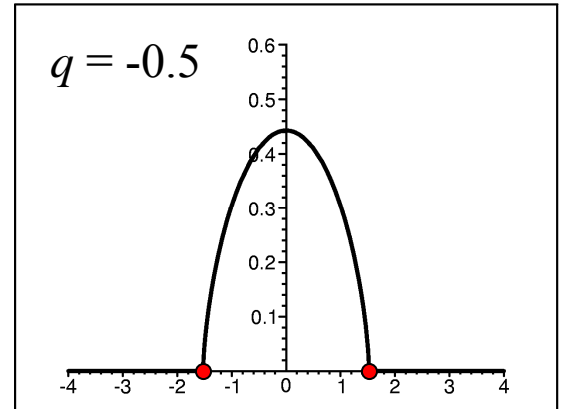
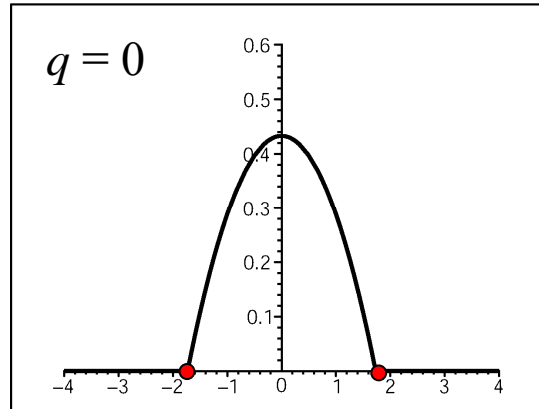
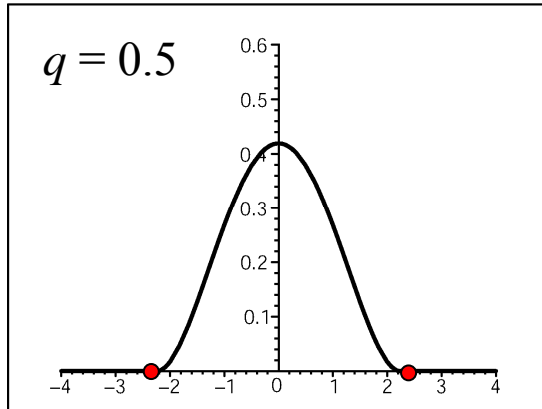
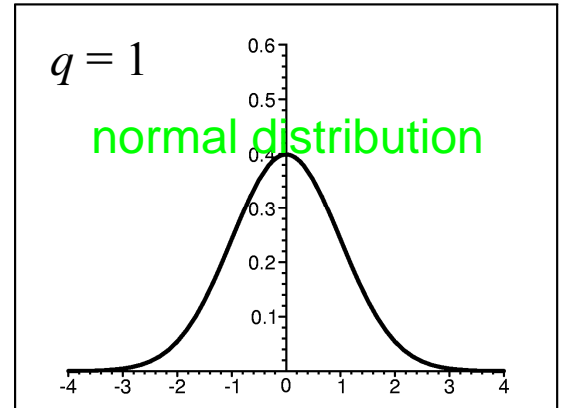
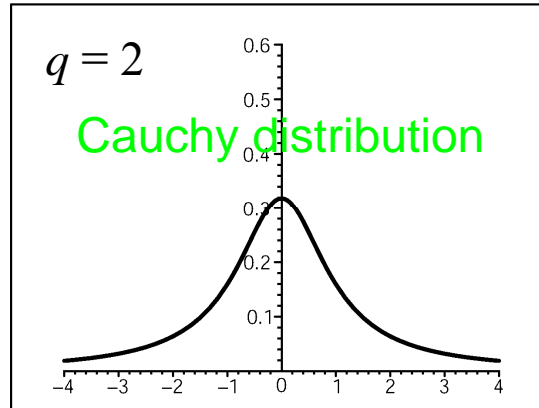
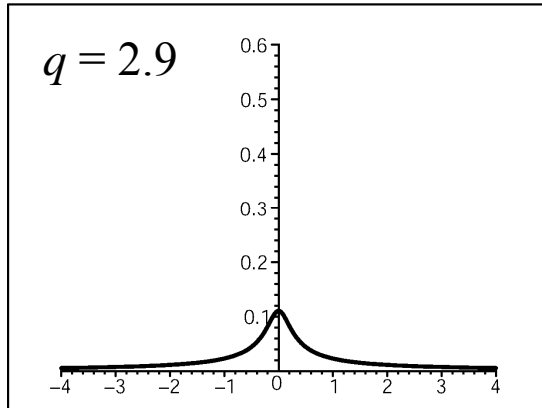
- for $q = 1 + \frac{2}{n+1}$, **t-distribution**

$$p_{1+\frac{2}{n+1}}(x) = \frac{1}{\sqrt{n}B\left(\frac{n}{2}, \frac{1}{2}\right)\sigma} \left\{1 + \frac{1}{n} \frac{(x-\mu)^2}{\sigma^2}\right\}^{-\frac{n+1}{2}}$$

If the covariance doesn't exist, σ should be interpreted as the scale factor.

The q relates the finiteness of the system and the dimension of the probability distribution.

q -normal distribution (The red points show cutoffs.)



NATURAL COORDINATE AND POTENTIAL FUNCTION

The q -normal distribution can be represented as

$$p_q(x) = \{1 + (1 - q)(\xi^1 x + \xi^2 x^2 - \psi_q)\}^{\frac{1}{1-q}} .$$

Natural coordinate system : $\left[\frac{\partial}{\partial \xi^1}, \frac{\partial}{\partial \xi^2} \right] = 0$ where $\begin{cases} \xi^1 = \Omega_q^2 \theta^1 \\ \xi^2 = \Omega_q^2 \theta^2 \end{cases}$ $\begin{cases} \theta^1 = \frac{\mu}{\sigma^2} \\ \theta^2 = -\frac{1}{2} \frac{1}{\sigma^2} \end{cases}$.

q -logarithmic likelihood : $L_q p_q = \frac{1}{1-q} (p_q^{1-q} - 1) = \xi^1 x + \xi^2 x^2 - \psi_q(\xi) ,$

Potential function : $\psi_q(\xi) = -\frac{1}{4} \frac{(\xi^1)^2}{\xi^2} - \frac{1}{1-q} \left\{ \frac{3-q}{2} \Omega_q^2 - 1 \right\}$

$$\begin{cases} \partial_1 \psi_q(\xi) = -\frac{1}{2} \frac{\xi^1}{\xi^2} = E_q[x] \\ \partial_2 \psi_q(\xi) = \frac{1}{4} \left(\frac{\xi^1}{\xi^2} \right)^2 - \Omega_q^2 \frac{1}{2\xi^2} = E_q[x^2] \end{cases} , \text{ where } \partial_i = \frac{\partial}{\partial \xi^i} .$$

METRIC TENSOR

$$\begin{cases} g_{ij} = \partial_i \partial_j \Psi_q(\xi) = -\partial_i \partial_j L_q p_q \text{ from the potential function} \\ g_{ij} = E_q[-\partial_i \partial_j L_q p_q] = q E_q \left[p_q^{\frac{q-1}{2}} (\partial_i L_q p_q) p_q^{\frac{q-1}{2}} (\partial_j L_q p_q) \right] \text{ from } q\text{-expectation} \end{cases}$$

$$\begin{aligned} g_{ij} &= q E_q \left[p_q^{\frac{q-1}{2}} (\partial_i L_q p_q) p_q^{\frac{q-1}{2}} (\partial_j L_q p_q) \right] \\ &= q \Omega_q^2 \int dx p_q^{-1} (\partial_i p_q) (\partial_j p_q) = \underbrace{q \Omega_q^2}_{\text{Scale factor}} E[\underbrace{(\partial_i \log p_q) (\partial_j \log p_q)}_{\text{Fisher metric}}] \end{aligned}$$



Non-Fisher metric !!

In the matrix form,

$$g_{ij} = \begin{pmatrix} -\frac{1}{2} \frac{1}{\xi^2} & \frac{1}{2} \frac{\xi^1}{(\xi^2)^2} \\ \frac{1}{2} \frac{\xi^1}{(\xi^2)^2} & -\frac{1}{2} \frac{(\xi^1)^2}{(\xi^2)^3} + \frac{1}{3-q} \Omega_q^2 \frac{1}{(\xi^2)^2} \end{pmatrix}.$$

The inverse of metric tensor :

$$g^{ij} = (3 - q) \Omega_q^{-2} \begin{pmatrix} (\xi^1)^2 - \frac{2}{3-q} \Omega_q^2 \xi^2 & \xi^1 \xi^2 \\ \xi^1 \xi^2 & (\xi^2)^2 \end{pmatrix}.$$

METRIC TENSOR AND DIVERGENCE

For the exponential family : $p(x|\xi) = \exp\{\xi^i X_i - \psi(\xi)\}$, the expectation : $\eta_i = E[X_i]$

$$\begin{aligned} D(p(x|\xi) || p(x|\xi + \rho)) &= \int dx p(x|\xi) \log \frac{p(x|\xi)}{p(x|\xi + \rho)} \\ &\simeq \frac{1}{2} \frac{\partial^2}{\partial \rho^i \partial \rho^j} D(p(x|\xi) || p(x|\xi + \rho)) \Big|_{\rho=0} \rho^i \rho^j \\ &= \frac{1}{2} \left\{ \int dx p(x|\xi) \frac{\partial \log p(x|\xi + \rho)}{\partial \rho^i} \frac{\partial \log p(x|\xi + \rho)}{\partial \rho^j} \Big|_{\rho=0} \right\} \rho^i \rho^j \\ &= \frac{1}{2} E \left[\frac{\partial \log p(x|\xi + \rho)}{\partial \rho^i} \frac{\partial \log p(x|\xi + \rho)}{\partial \rho^j} \Big|_{\rho=0} \right] \rho^i \rho^j = \frac{1}{2} g_{ij} \rho^i \rho^j \end{aligned}$$

$$\begin{aligned} D(p(x|\xi) || p(x|\xi + \rho)) &= \psi(\xi + \rho) - \psi(\xi) - \rho^i \eta_i \\ &\simeq \left(\frac{\partial \psi(\xi + \rho)}{\partial \rho^i} \Big|_{\rho=0} - \eta_i \right) \rho^i + \frac{1}{2} \frac{\partial^2 \psi(\xi + \rho)}{\partial \rho^i \partial \rho^j} \Big|_{\rho=0} \rho^i \rho^j \\ &= \frac{1}{2} \frac{\partial^2 \psi(\xi + \rho)}{\partial \rho^i \partial \rho^j} \Big|_{\rho=0} \rho^i \rho^j \end{aligned}$$



g_{ij} : positive definite and symmetric !!

However, for the case of q -normal distributions, ...

METRIC TENSOR AND q -DIVERGENCE

$$\frac{\partial L_q p_q(x|\xi)}{\partial \xi^i} = p(x|\xi)^{-q} \frac{\partial p_q(x|\xi)}{\partial \xi^i} \quad , \text{ the } q\text{-expectation : } \eta_i = E_q[X_i]$$

$$\begin{aligned} D_q(p_q(x|\xi) \| p_q(x|\xi + \rho)) &= \frac{\Omega_q^2(\xi)}{1-q} \left\{ 1 - \int dx p_q(x|\xi)^q p_q(x|\xi + \rho)^{1-q} \right\} \\ &\simeq \frac{1}{2} \frac{\partial^2}{\partial \rho^i \partial \rho^j} D_q(p_q(x|\xi) \| p_q(x|\xi + \rho)) \Big|_{\rho=0} \rho^i \rho^j \\ &= \frac{1}{2} \left\{ q \Omega_q^2(\xi) \int dx \frac{1}{p_q(x|\xi)} \frac{\partial p_q(x|\xi + \rho)}{\partial \rho^i} \frac{\partial p_q(x|\xi + \rho)}{\partial \rho^j} \Big|_{\rho=0} \right\} \rho^i \rho^j \\ &= \frac{1}{2} E_q \left[p_q(x|\xi)^{\frac{q-1}{2}} \frac{\partial L_q p_q(x|\xi + \rho)}{\partial \rho^i} p_q(x|\xi)^{\frac{q-1}{2}} \frac{\partial L_q p_q(x|\xi + \rho)}{\partial \rho^j} \Big|_{\rho=0} \right] \rho^i \rho^j \end{aligned}$$

$$\begin{aligned} D_q(p_q(x|\xi) \| p_q(x|\xi + \rho)) &= \psi_q(\xi + \rho) - \psi_q(\xi) - \rho^i \eta_i \\ &\simeq \left(\frac{\partial \psi_q(\xi + \rho)}{\partial \rho^i} \Big|_{\rho=0} - \eta_i \right) \rho^i + \frac{1}{2} \frac{\partial^2 \psi_q(\xi + \rho)}{\partial \rho^i \partial \rho^j} \Big|_{\rho=0} \rho^i \rho^j \\ &= \frac{1}{2} \frac{\partial^2 \psi_q(\xi + \rho)}{\partial \rho^i \partial \rho^j} \Big|_{\rho=0} \rho^i \rho^j \end{aligned}$$



g_{ij} : positive definite and symmetric but non-Fisher metric !!

DIVERGENCE AND q -PYTHAGOREAN THEOREM

q -divergence :

$$\begin{aligned}
 D_q(p(x|\xi^{(A)})|| p(x|\xi^{(B)})) &= \frac{\Omega_q^{(A)2}}{1-q} \left\{ 1 - \int dx p(x|\xi^{(A)})^q p(x|\xi^{(B)})^{1-q} \right\} \\
 &= q\Omega_q^{(A)2} D_{\alpha=1-2q}(p(x|\xi^{(A)})|| p(x|\xi^{(B)})) \\
 &= \varphi_q^{(A)} + \psi_q^{(B)} - \eta_i^{(A)} \xi^{(B)i}
 \end{aligned}$$

Legendre transformation : $\varphi_q = \eta_i \xi^i - \psi_q$ and $\eta_i = \frac{\partial \psi_q(\xi)}{\partial \xi^i}$

α -divergence :

$$D_\alpha(p(x|\xi^{(A)})|| p(x|\xi^{(B)})) = \frac{4}{1-\alpha^2} \left\{ 1 - \int dx p(x|\xi^{(A)})^{\frac{1-\alpha}{2}} p(x|\xi^{(B)})^{\frac{1+\alpha}{2}} \right\}$$

q -Pythagorean theorem :

$$\begin{aligned}
 &D_q(p(x|\xi^{(A)})|| p(x|\xi^{(B)})) + D_q(p(x|\xi^{(B)})|| p(x|\xi^{(C)})) - D_q(p(x|\xi^{(A)})|| p(x|\xi^{(C)})) \\
 &= (\xi^{(C)i} - \xi^{(B)i}) \left(\eta_i^{(A)} - \eta_i^{(B)} \right)
 \end{aligned}$$

PROBABILITY AMPLITUDE

Probability amplitude : $u_q = 2\sqrt{p_q}$.

The 1st order derivative with respect to the natural coordinate :

$$u_{q,i} = \partial_i(2\sqrt{p_q}) = p_q^{-\frac{1}{2}}(\partial_i p_q) = p_q^{\frac{q}{2}} \left\{ p_q^{\frac{q-1}{2}} (\partial_i L_q p_q) \right\} ,$$

$$\text{Identity : } \Omega_q^2 \int dx u_q u_{q,i} = 2\{E_q[x^i] - \partial_i \Psi_q(\xi)\} = 0 .$$

The 2nd order derivative with respect to the natural coordinates :

$$u_{q,ij} = p_q^{-\frac{1}{2}} \left\{ \left(q - \frac{1}{2} \right) p_q^{2q-1} (\partial_i L_q p_q) (\partial_j L_q p_q) + p_q^q (\partial_i \partial_j L_q p_q) \right\} .$$

$$\text{Metric tensor : } g_{ij} = q\Omega_q^2 \int dx u_{q,i} u_{q,j} = -q\Omega_q^2 \int dx u_q u_{q,ij} .$$

The 1st kind Christoffel symbol :

$$\begin{aligned} \Gamma_{kij} &= \frac{1}{2} \{ \partial_i g_{kj} + \partial_j g_{ik} - \partial_k g_{ij} \} = \frac{1}{2} \partial_i \partial_j \partial_k \Psi_q(\xi) \\ &= \left(q - \frac{1}{2} \right) q \Omega_q^2 \int dx p_q^{-\frac{1}{2}} u_{q,i} u_{q,j} u_{q,k} \\ &\quad + \frac{1}{2} \{ \Omega_q^{-2} \partial_k \Omega_q^2 g_{ij} + \Omega_q^{-2} \partial_j \Omega_q^2 g_{ik} + \Omega_q^{-2} \partial_i \Omega_q^2 g_{kj} \} . \end{aligned}$$

GAUGE THEORETIC DERIVATION OF α -CONNECTION

Local gauge transformation : $u_{q,i}^{(\alpha)} = e^{-\frac{\alpha}{2}L_q p_q} u_{q,i} .$

α -conjugation : $\alpha^\# = -\alpha , (-\alpha)^\# = \alpha .$

For example, $\left(u_{q,i}^{(\alpha)}\right)^\# = u_{q,i}^{(-\alpha)} = e^{\frac{\alpha}{2}L_q p_q} u_{q,i} .$

The local gauge transformation becomes U(1) under α -conjugation.

Metric tensor redefined : $g_{ij} = q\Omega_q^2 \int dx \left(u_{q,i}^{(\alpha)}\right)^\# u_{q,j}^{(\alpha)} = q\Omega_q^2 \int dx u_{q,i} u_{q,j} .$

Covariant derivative **before** local gauge transformation :

$$D_j u_{q,i} = \partial_j u_{q,i} - \Gamma_{ij}^\ell u_{q,\ell} .$$

Covariant derivative **after** local gauge transformation :

$$\begin{aligned} D_j^{(\alpha)} u_{q,i}^{(\alpha)} &= \partial_j u_{q,i}^{(\alpha)} - \Gamma_{ij}^{(\alpha)\ell} u_{q,\ell}^{(\alpha)} \\ &= e^{-\frac{\alpha}{2}L_q p_q} \left(u_{q,ij} - \Gamma_{ij}^{(\alpha)\ell} u_{q,\ell} - \frac{\alpha}{2} p_q^{-q+\frac{1}{2}} u_{q,j} u_{q,i} \right) . \end{aligned}$$

$$\left\{ \begin{array}{l} \text{Assumption : } \Gamma_{ij}^{(\alpha)\ell} = \Gamma_{ij}^{\ell} - \frac{\alpha}{2} A_{ij}^{\ell} \text{ under the local gauge transformation} \\ A_{ij}^{\ell} \text{ is called a gauge field.} \\ \text{Requirement : } D_j^{(\alpha)} u_{q,i}^{(\alpha)} = e^{-\frac{\alpha}{2} L_q p_q} D_j u_{q,i} \leftarrow \text{how to determine } A_{ij}^{\ell} \end{array} \right.$$

$$\text{From the assumption, } D_j^{(\alpha)} u_{q,i}^{(\alpha)} = e^{-\frac{\alpha}{2} L_q p_q} \left\{ D_j u_{q,i} + \frac{\alpha}{2} \left(A_{ij}^{\ell} u_{q,\ell} - p_q^{-q+\frac{1}{2}} u_{q,j} u_{q,i} \right) \right\} .$$

$$\begin{aligned} \text{Then } & q\Omega_q^2 \int dx \left(u_{q,k}^{(\alpha)} \right)^{\#} D_j^{(\alpha)} u_{q,i}^{(\alpha)} \\ & = q\Omega_q^2 \int dx u_{q,k} D_j u_{q,i} + \frac{\alpha}{2} \left(A_{ij}^{\ell} g_{\ell k} - q\Omega_q^2 \int dx p_q^{-q+\frac{1}{2}} u_{q,k} u_{q,j} u_{q,i} \right) . \end{aligned}$$

$$\text{From the requirement, } q\Omega_q^2 \int dx \left(u_{q,k}^{(\alpha)} \right)^{\#} D_j^{(\alpha)} u_{q,i}^{(\alpha)} = q\Omega_q^2 \int dx u_{q,k} D_j u_{q,i} .$$

$$\text{Therefore we have } A_{kij} = q\Omega_q^2 \int dx p_q^{-q+\frac{1}{2}} u_{q,k} u_{q,i} u_{q,j} , \text{ where } A_{kij} = A_{ij}^{\ell} g_{\ell k} .$$

Dualistic structure : $\Gamma_{kij}^{(\alpha)} + \Gamma_{kij}^{(-\alpha)} = 2\Gamma_{kij} = \partial_k g_{ij}$.

Covariant derivative of metric tensor with α -connection :

$$D_k^{(\alpha)} g_{ij} = D_k^{(\alpha)} \left\{ q\Omega_q^2 \int dx \left(u_{q,i}^{(\alpha)} \right)^\# u_{q,j}^{(\alpha)} \right\} = \alpha A_{ijk} .$$

Derived α -connection :

$$\Gamma_{111}^{(\alpha)} = 0$$

$$\Gamma_{112}^{(\alpha)} = (1 - \alpha\Omega_q^2) \frac{1}{4} \frac{1}{(\xi^2)^2}$$

$$\Gamma_{122}^{(\alpha)} = -(1 - \alpha\Omega_q^2) \frac{1}{2} \frac{\xi^1}{(\xi^2)^3}$$

$$\Gamma_{222}^{(\alpha)} = (1 - \alpha\Omega_q^2) \frac{3}{4} \frac{(\xi^1)^2}{(\xi^2)^4} - \left(1 - \alpha\Omega_q^2 \frac{3-q}{5-3q} \right) \Omega_q^2 \frac{5-q}{2(3-q)^2} \frac{1}{(\xi^2)^3} ,$$

α-RIEMANN TENSOR

Definition : $[D_k^{(\alpha)}, D_j^{(\alpha)}]U_i = R_{ijk}^{(\alpha)}U_\ell = g^{\ell m}R_{mijk}^{(\alpha)}U_\ell$,

α-Riemann tensor (the others are 0.) :

$$R_{1212}^{(\alpha)} = \frac{1}{4} \frac{1}{3-q} \frac{1}{(\xi^2)^3} + \frac{\alpha}{2} \Omega_q^2 \frac{(1-q)(5-2q)}{(3-q)(5-3q)} \frac{1}{(\xi^2)^3} - \frac{\alpha^2}{4} \Omega_q^2 \frac{q\Omega_q^2}{5-3q} \frac{1}{(\xi^2)^3}$$

$$R_{2112}^{(\alpha)} = -\frac{1}{4} \frac{1}{3-q} \frac{1}{(\xi^2)^3} + \frac{\alpha}{2} \Omega_q^2 \frac{(1-q)(5-2q)}{(3-q)(5-3q)} \frac{1}{(\xi^2)^3} + \frac{\alpha^2}{4} \Omega_q^2 \frac{q\Omega_q^2}{5-3q} \frac{1}{(\xi^2)^3}$$

$$R_{1221}^{(\alpha)} = -R_{1212}^{(\alpha)}$$

$$R_{2121}^{(\alpha)} = -R_{2112}^{(\alpha)}$$

$$R_{2212}^{(\alpha)} = -\alpha \Omega_q^2 \frac{(1-q)(5-2q)}{(3-q)(5-3q)} \frac{\xi^1}{(\xi^2)^4}$$

$$R_{2221}^{(\alpha)} = \alpha \Omega_q^2 \frac{(1-q)(5-2q)}{(3-q)(5-3q)} \frac{\xi^1}{(\xi^2)^4},$$

$$R_{1212}^{(\alpha)} \neq -R_{2112}^{(\alpha)}$$

$$R_{2212}^{(\alpha)} \neq 0$$

$$R_{2221}^{(\alpha)} \neq 0.$$

The α-Riemann tensor isn't antisymmetric, that is, the q-normal distribution is **not conjugate symmetric**.

α -RICCI TENSOR

Definition : $R_{ij}^{(\alpha)} = g^{kl} R_{likj}^{(\alpha)} = R_{ikj}^{(\alpha)}$.

α -Ricci tensor :

$$R_{11}^{(\alpha)} = \Omega_q^{-2} \frac{1}{4} \frac{1}{\xi^2} + \frac{\alpha}{4} \frac{(q-1)(5-q)}{5-3q} \frac{1}{\xi^2} - \frac{\alpha^2}{4} \Omega_q^2 \frac{q(3-q)}{5-3q} \frac{1}{\xi^2}$$

$$R_{12}^{(\alpha)} = -\Omega_q^{-2} \frac{1}{4} \frac{\xi^1}{(\xi^2)^2} - \frac{\alpha}{4} \frac{(q-1)(5-q)}{5-3q} \frac{\xi^1}{(\xi^2)^2} + \frac{\alpha^2}{4} \Omega_q^2 \frac{q(3-q)}{5-3q} \frac{\xi^1}{(\xi^2)^2}$$

$$R_{21}^{(\alpha)} = -\Omega_q^{-2} \frac{1}{4} \frac{\xi^1}{(\xi^2)^2} - \frac{\alpha}{4} \frac{(q-1)(5-q)}{5-3q} \frac{\xi^1}{(\xi^2)^2} + \frac{\alpha^2}{4} \Omega_q^2 \frac{q(3-q)}{5-3q} \frac{\xi^1}{(\xi^2)^2}$$

$$R_{22}^{(\alpha)} = \Omega_q^{-2} \frac{1}{4} \left\{ \frac{(\xi^1)^2}{(\xi^2)^3} - \Omega_q^2 \frac{2}{3-q} \frac{1}{(\xi^2)^2} \right\} \\ + \frac{\alpha}{4} \frac{(q-1)(5-q)}{5-3q} \left\{ \frac{(\xi^1)^2}{(\xi^2)^3} + \Omega_q^2 \frac{2}{3-q} \frac{1}{(\xi^2)^2} \right\} \\ - \frac{\alpha^2}{4} \Omega_q^2 \frac{q(3-q)}{5-3q} \left\{ \frac{(\xi^1)^2}{(\xi^2)^3} - \Omega_q^2 \frac{2}{3-q} \frac{1}{(\xi^2)^2} \right\} .$$

$$R_{12}^{(\alpha)} = R_{21}^{(\alpha)}$$

The α -Ricci tensor is symmetric, that is,
the q -normal distribution is **statistically equiaffine**.

The **Jeffreys' prior**, which plays an important role in statistical inference, is generalized to the **α -parallel prior** by J. Takeuchi and S. Amari (2005) from the information geometrical point of view.

However the α -parallel prior doesn't always exist for any value of α . So they give the sufficient condition for the existence of the α -parallel prior such that, if a statistical model is **conjugate symmetric**, the model is statistically **equiaffine**.

Then they leave **the open problem** such that whether there exists a statistically **equiaffine** natural model which is **not conjugate symmetric** or not.

➡ There exists such a model and the corresponding **α -parallel prior** is,

$$\exp \left\{ \alpha \frac{(5-2q)(3-q)}{5-3q} \Omega_q^2 S_q \right\} \sqrt{g} (d\xi^1 \wedge d\xi^2)$$

$$D_i^{(\alpha)} \left[\exp \left\{ \alpha \frac{(5-2q)(3-q)}{5-3q} \Omega_q^2 S_q \right\} \sqrt{g} (d\xi^1 \wedge d\xi^2) \right] = 0$$

THE JEFFREYS' PRIOR

In Bayesian statistical inference

Bayes rule for the posterior probability : $P(A|E) = \frac{P(A)}{P(E)} \int d^d \Theta \omega(\Theta) P(E|\Theta)$

A prior density on the parameter space : $\omega(\Theta)$, $A \in$ the model family

A collection of outcomes (i.i.d.) : E

Ignorance on the prior density : $\omega(\Theta)$

→ The uniform distribution on the parameter space

→ However, if a prior is uniform in the parameter space, the probability of a model family will depend on the arbitrary parametrization.

→ The uniform distribution NOT on parameters BUT distributions

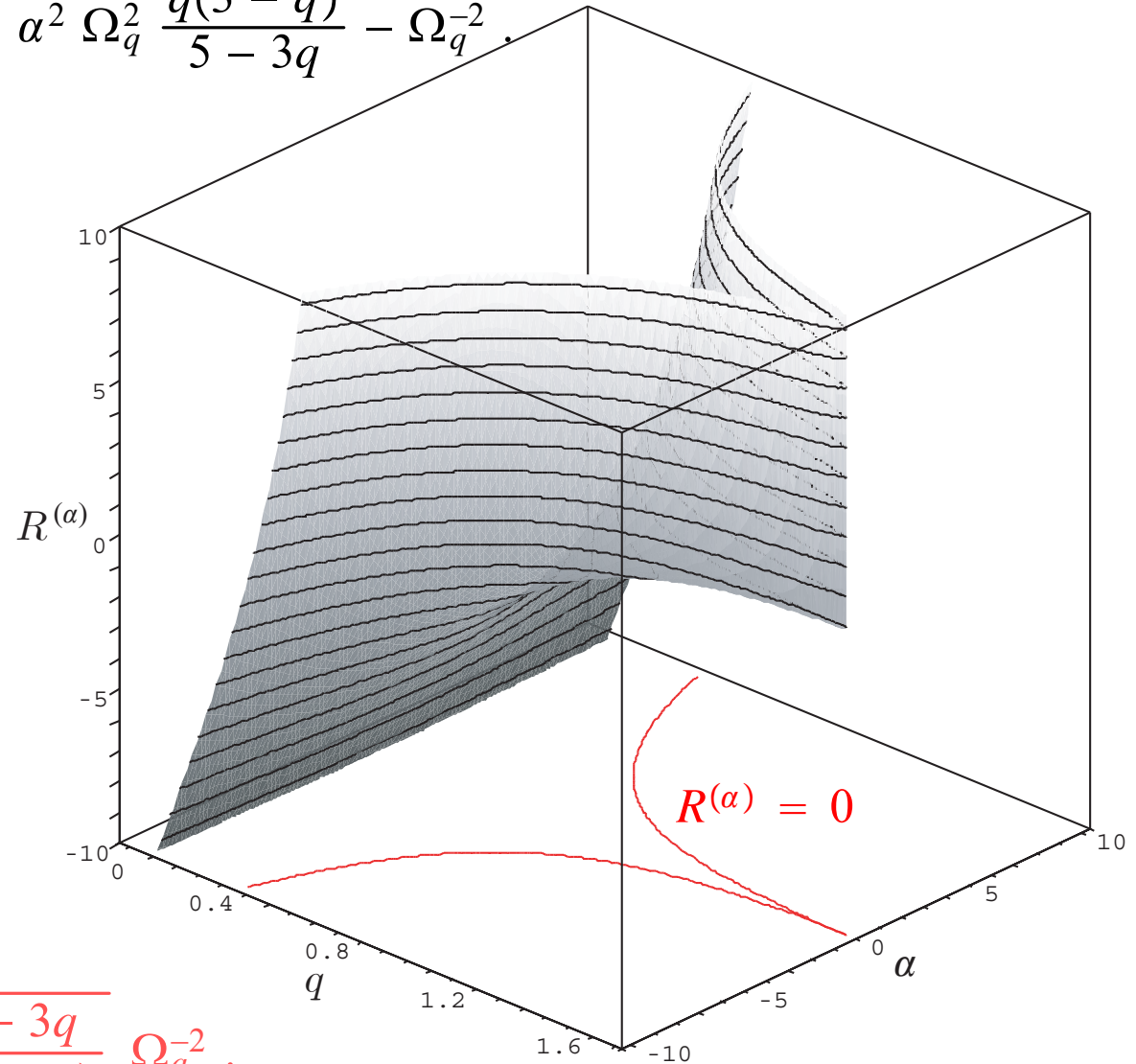
→ The invariant measure on the parameter space under reparametrizations

→ The Jeffreys' prior : $\sqrt{g} d^d \Theta$

α -RICCI SCALAR CURVATURE

α -Ricci scalar curvature :

$$R^{(\alpha)} = g^{ij}R_{ji}^{(\alpha)} = R_i^{(\alpha)i} = \alpha^2 \Omega_q^2 \frac{q(3-q)}{5-3q} - \Omega_q^{-2}.$$



$$R^{(\alpha)} = 0 \text{ at } \alpha = \pm \sqrt{\frac{5-3q}{q(3-q)}} \Omega_q^{-2}.$$

